Assimilation Algorithms: Ensemble Kalman Filters

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Outline

- The Standard Kalman Filter
- Kalman Filters for large dimensional systems
- Approximate Kalman Filters: the Ensemble Kalman Filter
- Ensemble Kalman Filters in hybrid Data Assimilation

Standard Kalman Filter

- In the Overview of Assimilation Methods lecture we have seen that, assuming all errors have Gaussian statistic, the posterior (i.e., analysis) distribution $p(\boldsymbol{x}|\boldsymbol{y})$ can also be expressed as a Gaussian distribution:

$$p(\boldsymbol{y}|\boldsymbol{x}) = \frac{1}{(2\pi)^{N/2}|\mathbf{R}|^{1/2}} \boldsymbol{exp}\left(-\frac{1}{2}\left(\boldsymbol{y}-\boldsymbol{H}(\boldsymbol{x})\right)^{\boldsymbol{T}}(\mathbf{R})^{-\mathbf{1}}\left(\boldsymbol{y}-\boldsymbol{H}(\boldsymbol{x})\right)\right)$$

$$p(\boldsymbol{x}_b|\boldsymbol{x}) = \frac{1}{(2\pi)^{N/2}|\mathbf{P}_B|^{1/2}} \boldsymbol{exp}\left(-\frac{1}{2}(\boldsymbol{x_b}-\boldsymbol{x})^{\boldsymbol{T}}(\mathbf{P}_B)^{-\mathbf{1}}(\boldsymbol{x_b}-\boldsymbol{x})\right)$$

$$p(\boldsymbol{x}|\boldsymbol{y}) \propto p(\boldsymbol{y}|\boldsymbol{x})p(\boldsymbol{x}_b|\boldsymbol{x}) \propto \boldsymbol{exp}\left(-\frac{1}{2}\left(\boldsymbol{y}-\boldsymbol{H}(\boldsymbol{x})\right)^{\boldsymbol{T}}(\mathbf{R})^{-\mathbf{1}}\left(\boldsymbol{y}-\boldsymbol{H}(\boldsymbol{x})\right)-\frac{1}{2}(\boldsymbol{x_b}-\boldsymbol{x})^{\boldsymbol{T}}(\mathbf{P}_B)^{-\mathbf{1}}(\boldsymbol{x_b}-\boldsymbol{x})\right)$$

- Kalman Filter methods try to find the mean and covariance of this posterior distribution

- Note that, under this Gaussian assumption, knowing the mean and covariance of $p(\boldsymbol{x}|\boldsymbol{y})$ means knowing the full posterior pdf

Standard Kalman Filter

- Let us start with a simple univariate example:

  Assume we are analysing a single state variable x, whose errors are zero mean and Gaussian distributed around its background forecast x_b:

$$p(x_b|x) = \frac{1}{(2\pi\sigma_b^2)^{1/2}} exp\left(-\frac{1}{2}\frac{(x_b - x)^2}{\sigma_b^2}\right) \sim \mathcal{N}(x_b, \sigma_b^2)$$

  We have one observation of the state variable, also with Gaussian errors:

$$p(y|x) = \frac{1}{(2\pi\sigma_o^2)^{1/2}} exp\left(-\frac{1}{2}\frac{(y - x)^2}{\sigma_o^2}\right) \sim \mathcal{N}(y, \sigma_o^2)$$

  Applying Bayes theorem we find:

$$p(x|y) \propto p(y|x)p(x_b|x) \propto exp\left(-\frac{1}{2}\frac{(y-x)^2}{\sigma_b^2} - \frac{1}{2}\frac{(x_b-x)^2}{\sigma_b^2}\right) \propto exp\left(-\frac{1}{2}\left(\left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2}\right)x^2 - 2\left(\frac{x_b}{\sigma_b^2} + \frac{y}{\sigma_o^2}\right)x\right)\right)$$

  Comparing to a standard Gaussian distribution:

$$p(x) = \frac{1}{(2\pi\sigma^2)^{1/2}} exp\left(-\frac{1}{2}\frac{(x-\mu)^2}{\sigma^2}\right) = exp\left(-\frac{1}{2}\left(\left(\frac{1}{\sigma^2}\right)x^2 - 2\left(\frac{\mu}{\sigma^2}\right)x + \left(\frac{\mu^2}{\sigma^2}\right)\right)\right)$$

  we see that the posterior distribution is also Gaussian with mean and variance:

$$Var(x|y) = \sigma^2 = \left(\frac{1}{\sigma_b^2} + \frac{1}{\sigma_o^2}\right)^{-1} = \frac{\sigma_o^2\sigma_b^2}{\sigma_o^2 + \sigma_b^2}$$

$$E(x|y) = \mu = \sigma^2\left(\frac{x_b}{\sigma_b^2} + \frac{y}{\sigma_o^2}\right) = \frac{\sigma_o^2\sigma_b^2}{\sigma_o^2 + \sigma_b^2}\left(\frac{x_b}{\sigma_b^2} + \frac{y}{\sigma_o^2}\right) = \frac{\sigma_o^2}{\sigma_o^2 + \sigma_b^2}x_b + \frac{\sigma_b^2}{\sigma_o^2 + \sigma_b^2}\,y$$

Standard Kalman Filter

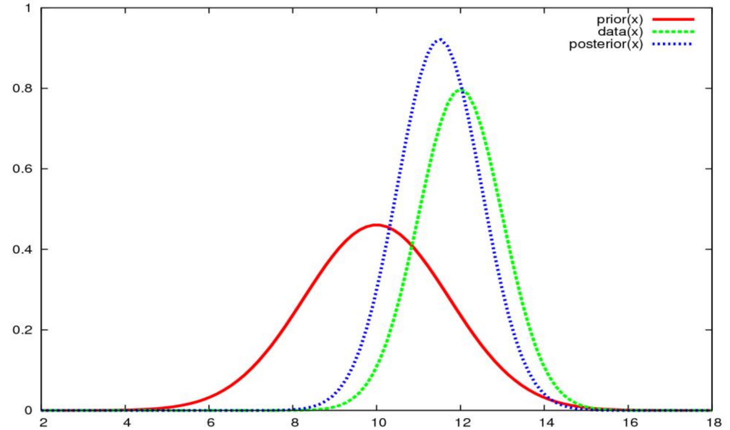
- A simple univariate example:

Introducing the Kalman gain: $K = \frac{\sigma_b^2}{\sigma_o^2 + \sigma_b^2}$ the equations for the mean and variance can be recast as:

$$Var(x|y) = \frac{\sigma_o^2 \sigma_b^2}{\sigma_o^2 + \sigma_b^2} = (1-K)\sigma_b^2$$

$$E(x|y) = \frac{\sigma_o^2}{\sigma_o^2 + \sigma_b^2} x_b + \frac{\sigma_b^2}{\sigma_o^2 + \sigma_b^2} \mathrm{y} = x_b + K(y - x_b)$$

The posterior variance is thus reduced (1-K<0) with respect to the prior (background) variance, while the posterior mean is a linear, weighted average of the prior and the observation.



Standard Kalman Filter

- These Kalman Filter analysis update equations can be generalised to the multi-dimensional and multivariate case (Wikle and Berliner, 2007; Bromiley, 2014):

$$E(\boldsymbol{x}|\boldsymbol{y}) = \boldsymbol{x_a} = \boldsymbol{x_b} + \mathbf{K}(\boldsymbol{y} - \mathbf{H}(\boldsymbol{x_b}))$$

$$\begin{aligned} Var(\boldsymbol{x}|\boldsymbol{y}) = \mathbf{P}^a = \; & (\mathbf{I} - \mathbf{KH})\mathbf{P}^b(\mathbf{I} - \mathbf{KH})^T + \mathbf{KRK}^{\boldsymbol{T}} = \\ & (\mathbf{I} - \mathbf{KH})\mathbf{P}^b - \mathbf{P}^b\mathbf{H}^T\mathbf{K}^T + \mathbf{K}\left(\mathbf{HP}^b\mathbf{H}^T + \mathbf{R}\right)\mathbf{K}^T = \\ & (\mathbf{I} - \mathbf{KH})\mathbf{P}^b \end{aligned}$$

$$\mathbf{K} = \mathbf{P}^b\mathbf{H}^T\left(\mathbf{HP}^b\mathbf{H}^T + \mathbf{R}\right)^{-1} = \left(\left(\mathbf{P}^b\right)^{-1} + \mathbf{H}^T\mathbf{R}^{-1}\mathbf{H}\right)^{-1}\mathbf{H}^T\mathbf{R}^{-1}$$

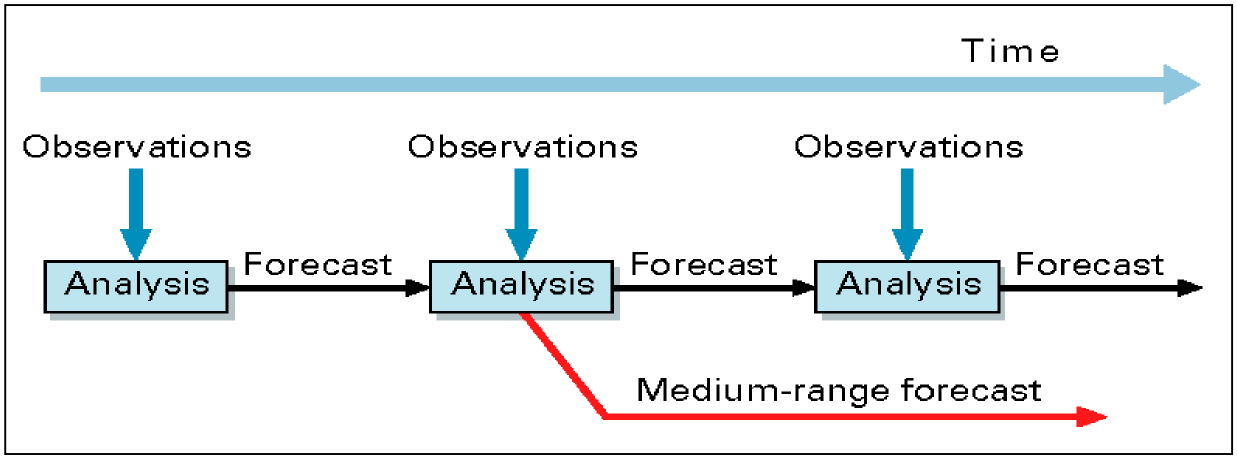
These are the same update equations obtained in Assimilation Algorithms (1).

In that context they were derived without making assumptions of Gaussianity, but looking for the analysis estimate which had the minimum error variance and could be expressed as a linear combination of the background and the observations (we called it the BLUE, Best Linear Unbiased Estimate). Linearity of observation operator and model was invoked.

If the background and observations are normally distributed, the Kalman Filter update equations give us the mean and the covariance of the posterior distribution. Under these hypotheses the posterior distribution is also Gaussian, so we have completely solved the problem!

Standard Kalman Filter

- In NWP applications of data assimilation we want to update our estimate of the state and its uncertainty at later times, as new observations come in: we want to cycle the analysis



- For each analysis in this cycle we require a background $\mathbf{x}^b{}_t$ (i.e. a prior estimate of the state valid at time t)

- Usually, our best prior estimate of the state at time t is given by a forecast from the preceding analysis at time t-1 (the "background"):

$$\mathbf{x}^b{}_t = \mathbf{M}(\mathbf{x}^a{}_{t-1})$$

- What is the error covariance matrix (=> the uncertainty) associated with this background?

Standard Kalman Filter

- What is the error covariance matrix associated with the background forecast?

$$\mathbf{x}^b{}_t = \mathbf{M}(\mathbf{x}^a{}_{t-1})$$

- Subtract the true state $\mathbf{x}^*$ from both sides of the equation:

$$\mathbf{x}^b - \mathbf{x}^*{}_t = \boldsymbol{\varepsilon}^b{}_t = \mathbf{M}(\mathbf{x}^a{}_{t-1}) - \mathbf{x}^*{}_t$$

- Since $\mathbf{x}^a{}_{t-1} = \mathbf{x}^*{}_{t-1} + \boldsymbol{\varepsilon}^a{}_{t-1}$ we have:

$$\boldsymbol{\varepsilon}^b{}_t = \mathbf{M}(\mathbf{x}^*{}_{t-1} + \boldsymbol{\varepsilon}^a{}_{t-1}) - \mathbf{x}^*{}_t \approx$$

$$\mathbf{M}(\mathbf{x}^*{}_{t-1}) + \mathbf{M}\boldsymbol{\varepsilon}^a{}_{t-1} - \mathbf{x}^*{}_t =$$

$$\mathbf{M}\boldsymbol{\varepsilon}^a{}_{t-1} + \boldsymbol{\eta}_t$$

- Here we have defined the model error $\boldsymbol{\eta}_t = \mathbf{M}(\mathbf{x}^*{}_{t-1}) - \mathbf{x}^*{}_t$

- We will also assume that no systematic errors are present in our system

$$< \boldsymbol{\varepsilon}^a > = < \boldsymbol{\eta} > = 0 \Rightarrow < \boldsymbol{\varepsilon}^b > = 0$$

Standard Kalman Filter

- The background error covariance matrix will then be given by:

$$<\boldsymbol{\varepsilon}^b_t (\boldsymbol{\varepsilon}^b_t)^T> \triangleq \mathbf{P}^b_t = <(\mathbf{M}\boldsymbol{\varepsilon}^a_{t-1} + \eta_t)\ (\mathbf{M}\boldsymbol{\varepsilon}^a_{t-1} + \eta_t)^T> =$$

$$\mathbf{M}<\boldsymbol{\varepsilon}^a_{t-1} (\boldsymbol{\varepsilon}^a_{t-1})^T> \mathbf{M}^T + <\boldsymbol{\eta}_t (\boldsymbol{\eta}_t)^T> =$$

$$\mathbf{M}\,\mathbf{P}^a_{t-1}\,\mathbf{M}^T + \mathbf{Q}_t$$

- Here we have assumed $< \boldsymbol{\varepsilon}^a_{t-1} (\boldsymbol{\eta}_t)^T> = 0$ and defined the model error covariance matrix $\mathbf{Q}_t \triangleq <\boldsymbol{\eta}_t (\boldsymbol{\eta}_t)^T>$

- Note how the background error is described as the sum of the errors of the previous analysis propagated by the model dynamics to the time of the update ($\mathbf{M}\,\mathbf{P}^a_{t-1}\,\mathbf{M}^T$) and the new errors introduced by the model integration ($\mathbf{Q}_t$)

- We now have all the equations necessary to propagate and update both the state and its error estimates

Standard Kalman Filter

$\mathbf{x}^a_{t-1}$
$\mathbf{P}^a_{t-1}$

t-1 → t → t+1

New Observations

Propagation (t-1 → t)

1. Predict the state ahead
   $\mathbf{x}^b_t = \mathbf{M}(\mathbf{x}^a_{t-1})$
2. Predict the state error cov.
   $\mathbf{P}^b_t = \mathbf{M}\,\mathbf{P}^a_{t-1}\,\mathbf{M}^T + \mathbf{Q}_t$

Update

3. Compute the Kalman Gain
   $\mathbf{K} = \mathbf{P}^b\,\mathbf{H}^T(\mathbf{H}\,\mathbf{P}^b\,\mathbf{H}^T + \mathbf{R})^{-1}$
4. Update state estimate
   $\mathbf{x}^a_t = \mathbf{x}^b_t + \mathbf{K}\,(\mathbf{y} - \mathbf{H}(\mathbf{x}^b_t))$
5. Update state error estimate
   $\mathbf{P}^a_t = (\mathbf{I} - \mathbf{KH})\mathbf{P}^b_t\,(\mathbf{I} - \mathbf{KH})^T + \mathbf{KRK}^T$

Propagation (t → t+1)

1. Predict the state ahead
   $\mathbf{x}^b_{t+1} = \mathbf{M}(\mathbf{x}^a_t)$
2. Predict the state error cov.
   $\mathbf{P}^b_{t+1} = \mathbf{M}\,\mathbf{P}^a_t\,\mathbf{M}^T + \mathbf{Q}_{t+1}$

Standard Kalman Filter

- Under the assumption that the model $\mathbf{M}$ and the observation operator $\mathbf{H}$ are linear operators (i.e., they do not depend on the state $\mathbf{x}$), the Kalman Filter produces an optimal sequence of analyses ($\mathbf{x}^a_1$, $\mathbf{x}^a_2$, …, $\mathbf{x}^a_{t-1}$, $\mathbf{x}^a_t$)

- The KF analysis $\mathbf{x}^a_t$ is the best (minimum variance) linear unbiased estimate of the state at time t, given $\mathbf{x}^b_t$ and all observations up to time t ($\mathbf{y}_0$,$\mathbf{y}_1$,…,$\mathbf{y}_t$).

- Note that Gaussianity of errors is not required. If errors are Gaussian the Kalman Filter provides the correct posterior conditional probability estimate (according to Bayes' Law), i.e. $p(\mathbf{x}^a_t | \mathbf{x}^b_0; \mathbf{y}_0,\mathbf{y}_1,\ldots,\mathbf{y}_t)$. This also implies that if errors are Gaussian then the state estimated with the KF is also the most likely state (the mode of the pdf).

Extended Standard Kalman Filter

- The extended Kalman Filter (EKF) is a non-linear extension of the Kalman Filter where the model and observation operators are not required to be linear operators (independent of the state) as in the standard KF:

$$\mathbf{y} = \mathcal{H}(\mathbf{x}_b) + \boldsymbol{\varepsilon}_o \quad \text{(EKF)} \qquad \mathbf{y} = \mathbf{H}\mathbf{x}_b + \boldsymbol{\varepsilon}_o \quad \text{(KF)}$$

$$\mathbf{x}_\mathrm{b} = \mathcal{M}(\mathbf{x}_a) + \boldsymbol{\eta} \quad \text{(EKF)} \qquad \mathbf{x}_b = \mathbf{M}\mathbf{x}_a + \boldsymbol{\eta} \quad \text{(KF)}$$

- The covariance update and prediction steps of the KF equations use the Jacobians of the model and observation operators, linearized around the analysed/predicted state, i.e.:

$$\mathbf{M} = \frac{\partial \mathcal{M}}{\partial \boldsymbol{x}}(\mathbf{x}_{\mathrm{a,t-1}}), \quad \mathbf{H} = \frac{\partial \mathcal{H}}{\partial \boldsymbol{x}}(\mathbf{x}_{\mathrm{b,t}})$$

- The EKF is thus a first order linearization of the KF equations around the current state estimates. As such it is as good as a first order Taylor expansion is a good approximation of the nonlinear system we are dealing with.

- A type of EKF is used at ECMWF in the analysis of soil variables (Simplified Extended Kalman Filter, SEKF). More on this later in the dedicated lecture.

Kalman Filters for large dimensional systems

- The Kalman Filter (standard or extended) is unfeasible for large dimensional systems

- The size N of the analysis/background state in the ECMWF 4DVar is $O(10^8)$: the KF requires us to store and evolve in time state covariance matrices ($\mathbf{P}^{a/b}$) of $O(NxN)$
  - The World's fastest computers can sustain ~ 10^{15} operations per second
  - An efficient implementation of matrix multiplication of two 10^8x10^8 matrices requires ~10^{22} ($O(N^{2.8})$)operations: about 4 months on the fastest computer!
  - Evaluating $\mathbf{P}^b{}_t = \mathbf{M}\,\mathbf{P}^a{}_{t-1}\,\mathbf{M}^T + \mathbf{Q}_k$ requires 2*N≈2*10^8 model integrations!

- A range of approximate Kalman Filters has been developed for use with large-dimensional systems.

- All of these methods rely on some forms of low-rank approximation of the state covariance matrices of background and analysis errors.

Kalman Filters for large dimensional systems

- Let us assume that $\mathbf{P}^{a/b}$ has rank M<<N (e.g. M≈100). (rank=dim. of vector space spanned by its columns/rows)
- Then we can write $\mathbf{P}^b = \mathbf{X}^b(\mathbf{X}^b)^T$, where $\mathbf{X}^b{}_k$ is N x M. This decomposition also assures us that the resulting $\mathbf{P}^b$ is positive definite.
- The Kalman Gain then becomes:

$$\mathbf{K} = \mathbf{P}^b\,\mathbf{H}^T(\mathbf{H}\,\mathbf{P}^b\,\mathbf{H}^T + \mathbf{R})^{-1} =$$

$$\mathbf{X}^b(\mathbf{X}^b)^T\mathbf{H}^T(\mathbf{H}\,\mathbf{X}^b(\mathbf{X}^b)^T\,\mathbf{H}^T + \mathbf{R})^{-1} =$$

$$\mathbf{X}^b\,(\mathbf{H}\mathbf{X}^b)^T(\mathbf{H}\,\mathbf{X}^b(\mathbf{H}\mathbf{X}^b)^T + \mathbf{R})^{-1}$$

- Note that, to evaluate $\mathbf{K}$, we apply $\mathbf{H}$ to the M columns of $\mathbf{X}^b$ rather than to the N columns of $\mathbf{P}^b$!!
- The N x N matrices $\mathbf{P}^{a/b}$ have been eliminated from the computation! In their place we have to deal with N x M ($\mathbf{X}^b$) matrices in state space and their observation space projections L x M ($\mathbf{H}\mathbf{X}^b$) matrices (L = number of observations)

Kalman Filters for large dimensional systems

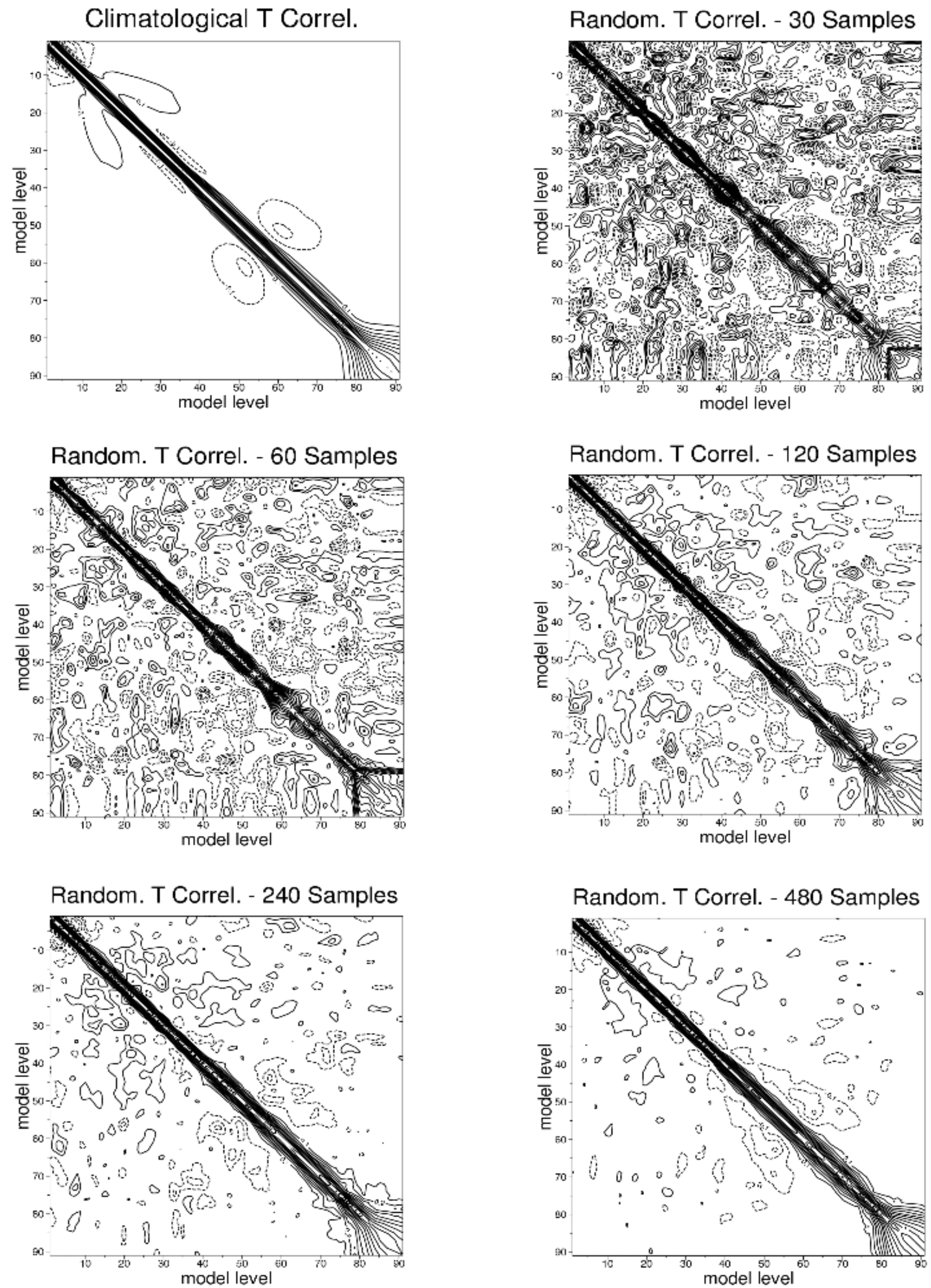
- The approximated KF described above is called Reduced-Rank Kalman Filter (RRKF)
- There is a price to pay for this huge gain in computational cost
- The analysis increment is a linear combination of the columns of $\mathbf{X}^b$:

$$\mathbf{x}^a - \mathbf{x}^b = \mathbf{K}\,(\mathbf{y} - \mathrm{H}(\mathbf{x}^b)) = \mathbf{X}^b\,(\mathbf{H}\mathbf{X}^b)^T\,((\mathbf{H}\mathbf{X}^b)(\mathbf{H}\mathbf{X}^b)^T + \mathbf{R})^{-1}\,(\mathbf{y} - \mathrm{H}(\mathbf{x}^b))$$

- The whole blue part of the equation computes to a vector of size M!
- The analysis increments are thus formed as a linear combination of the columns of $\mathbf{X}^b$: they are confined to the column subspace of $\mathbf{X}^b$, which has at most rank $M \ll N$.
- This severe reduction in rank of $\mathrm{P}^{a/b}$ has two main effects:
  1. There are too few degrees of freedom available to fit the $\approx 10^7$ observations available during the analysis window: the analysis is too “smooth”;
  2. The low-rank approximations of the covariance matrices suffer from spurious long-distance correlations.

Random sampling of vertical background error correlation matrix for different ensemble sizes.

Note how sampling noise decreases slowly with ensemble size $O(M^{1/2})$



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Kalman Filters for large dimensional systems

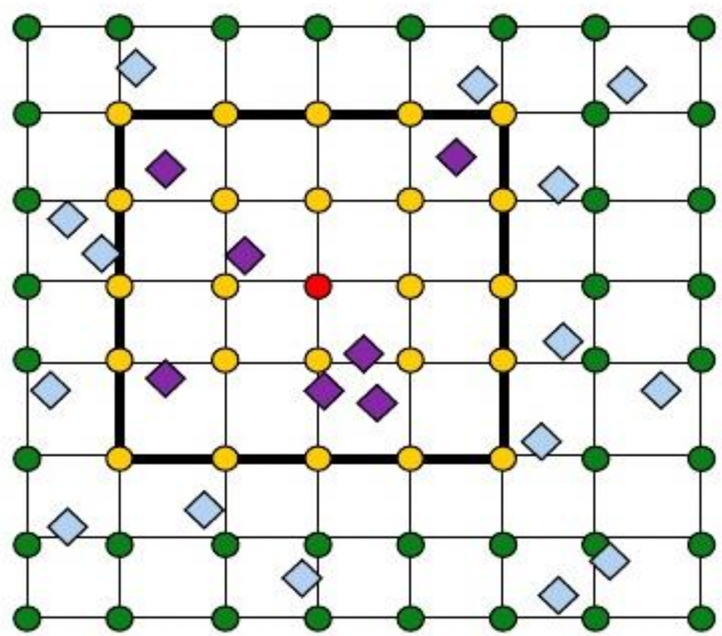
- There are two main ways to combat the rank deficiency/sampling noise problem:
    1. Domain localization (e.g. Houtekamer and Mitchell, 1998; Ott *et al.* 2004);
- Domain localization solves the analysis equations independently for each grid point, or for each of a set of regions covering the domain.
- Each analysis uses only observations that are local to the grid point (or region) and the observations are usually weighted according to their distance from the analysed grid point (e.g., Hunt *et al.*, 2007)
- This guarantees that the analysis at each grid point (or region) is not influenced by distant observations.
- The method acts to vastly increase the dimension of the sub-space in which the analysis increment is constructed because each grid point is updated by a different linear combination of ensemble perturbations
- However, performing independent analyses for each region can lead to difficulties in the analysis of the large scales and in producing balanced analyses.

Kalman Filters for large dimensional systems

- There are two ways around the rank deficiency problem:
  1. Domain localization (e.g. Houtekamer and Mitchell, 1998; Ott *et al.* 2004);

● Analysed grid point

◆ Local observations



Kalman Filters for large dimensional systems

2. Covariance localization (e.g. Houtekamer and Mitchell, 2001).

- Covariance localization is performed by element wise (Schur) multiplication of the error covariance matrices with a predefined correlation matrix representing a decaying function of distance (vertical and/or horizontal).

$$\mathbf{P}^b \dashrightarrow \rho_L \circ \mathbf{P}^b$$

- In this way spurious long range correlations in $\mathbf{P}^b$ are suppressed.
- As for domain localization, the method acts to vastly increase the dimension of the sub-space in which the analysis increment is constructed.
- Choosing the product function is non-trivial. It is easy to modify $\mathbf{P}^b$ in undesirable ways. In particular, balance relationships (e.g. geostrophy) may be adversely affected.
- In order to suppress sampling noise some of the information content of the observations is lost

Kalman Filters for large dimensional systems

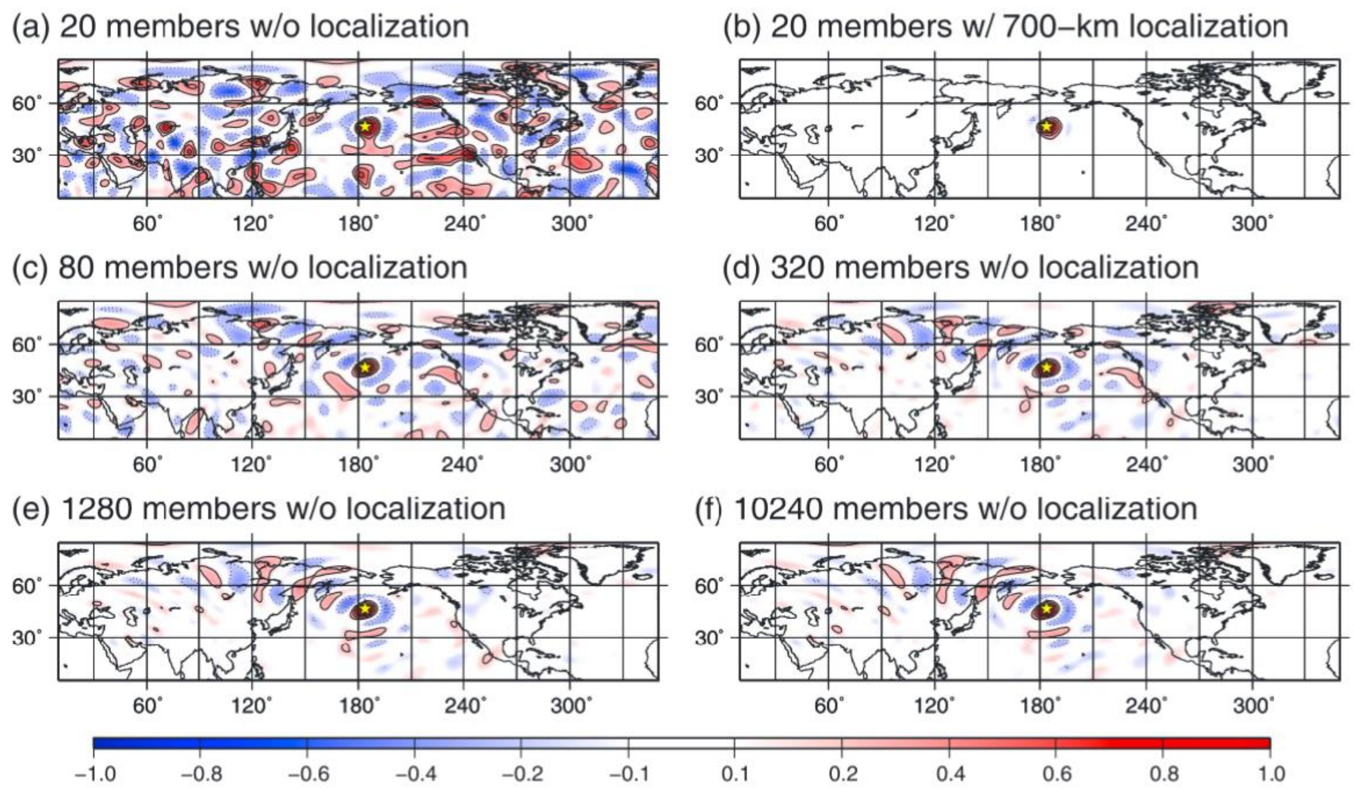
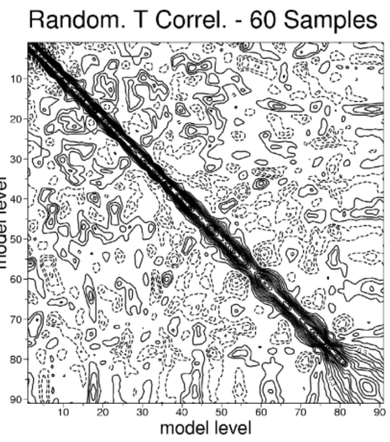


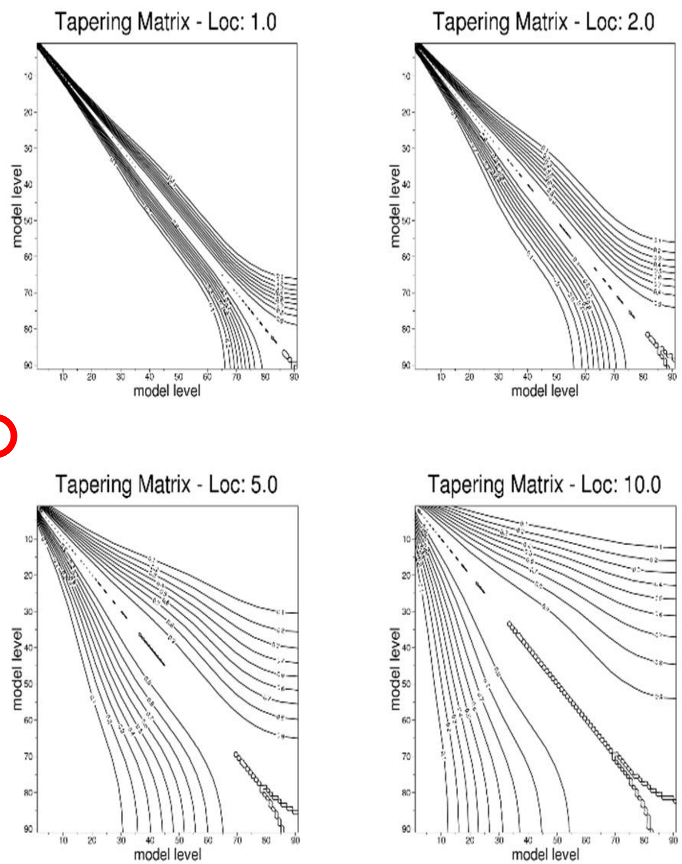
Figure 4. Similar to Figure 1 but at 00:00 UTC 18 January with the yellow star point at 46.389°N, 176.25°W and for different ensemble sizes ((a) 20, (c) 80, (d) 320, (e) 1280, and (f) 10,240 members) and (b) with localization for 20 members.

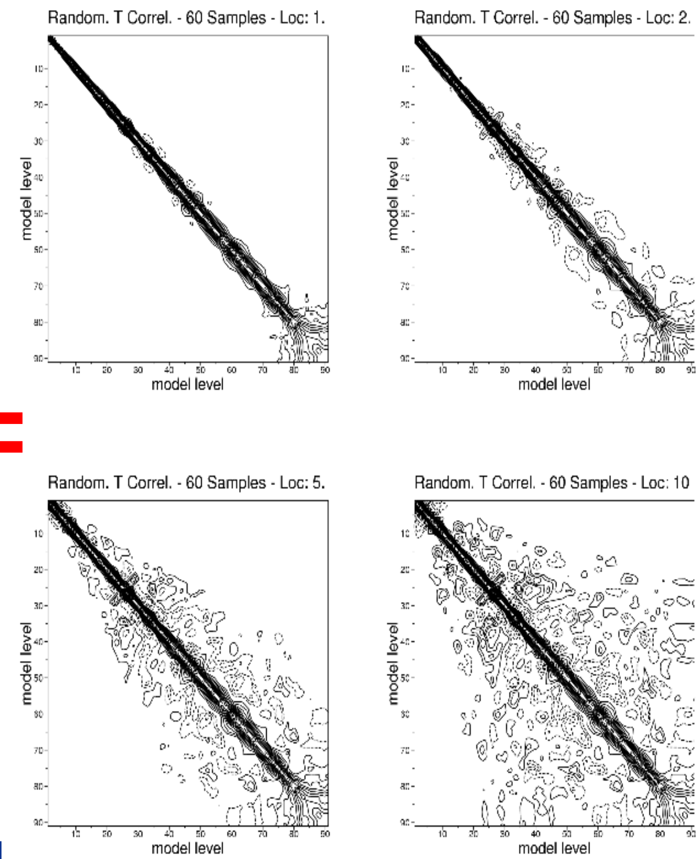
Miyoshi et al., 2014

- Standard Error of sample correlation ≈ $(1-\rho^2)/\sqrt{(N_{ens}-1)}$
- For small ρ, N_{ens} it becomes >= ρ (e.g. $\rho=0.1$, $N_{ens}=40$ => stderr(ρ)≈0.16)
- Since ρ -> 0 for large horiz./vert. distances apply distance based covariance localization on the sample $\mathbf{P}^f$

$P^f_{sampled} \circ \rho_L = P^f_{local}$







Kalman Filters for large dimensional systems

- Domain/Covariance localization is a practical necessity for using the KF in large dimensional applications
- Finding the right amount of localization is an (expensive) tuning exercise: a good trade-off needs to be found between computational effort, sampling error and imbalance error
- Finding the “optimal” localization scales as functions of the system characteristics is an area of active research (e.g., Flowerdew, 2015; Periáñez et al., 2014; Menetrier et al., 2014; Bishop, 2017)

Ensemble Kalman Filters

- Ensemble Kalman Filters (EnKF, Evensen, 1994; Houtekamer and Mitchell, 1998; Burgers et al., 1998) are Monte Carlo implementations of the reduced rank KF
- In EnKF error covariances are constructed as sample covariances from an ensemble of background/analysis fields, of size M:

$$\mathbf{P}^{\mathrm{b}} = \frac{1}{M-1}\Sigma_{\mathrm{m}}(\mathbf{x}^{\mathrm{b}}{}_{\mathrm{m}} - <\mathbf{x}^{\mathrm{b}}{}_{\mathrm{m}}>)\ (\mathbf{x}^{\mathrm{b}}{}_{\mathrm{m}} - <\mathbf{x}^{\mathrm{b}}{}_{\mathrm{m}}>)^{\mathrm{T}} =$$

$$= \mathbf{X}^{\mathrm{b}}(\mathbf{X}^{\mathrm{b}})^{\mathrm{T}}$$

- $\mathbf{X}^{\mathrm{b}}$ is the *N* x *M* matrix of normalised background perturbations, i.e.:

$$\mathbf{X}^{\mathrm{b}} = \frac{1}{\sqrt{M-1}}\left((\mathbf{x}^{\mathrm{b}}{}_{1} - <\mathbf{x}^{\mathrm{b}}>), (\mathbf{x}^{\mathrm{b}}{}_{2} - <\mathbf{x}^{\mathrm{b}}>), .., (\mathbf{x}^{\mathrm{b}}{}_{M} - <\mathbf{x}^{\mathrm{b}}>)\right)$$

- Note that the full covariance matrix is never formed explicitly: The error covariances are usually computed locally for each grid point in the *M* x *M* ensemble space

Ensemble Kalman Filters

- In the standard KF the error covariances are explicitly computed and propagated in time using the tangent linear and adjoint of the model and observation operators, i.e.:

$$\mathbf{K} = \mathbf{P}^b \mathbf{H}^T(\mathbf{H}\mathbf{P}^b\mathbf{H}^T + \mathbf{R})^{-1}$$

$$\mathbf{P}^b = \mathbf{M}\mathbf{P}^a\mathbf{M}^T + \mathbf{Q}$$

- In the EnKF the error covariances are sampled from the ensemble forecasts and the huge matrix $\mathbf{P}^b$ is never explicitly formed:

$$\mathbf{P}^b\mathbf{H}^T = \mathbf{X}^b(\mathbf{X}^b)^T\mathbf{H}^T = \mathbf{X}^b(\mathbf{H}\mathbf{X}^b)^T =$$

$$\frac{1}{M-1}\Sigma_m(\mathbf{x}^b{}_m - <\mathbf{x}^b{}_m>)\ (\mathbf{H}\mathbf{x}^b{}_m - <\mathbf{H}(\mathbf{x}^b{}_m)>)^T$$

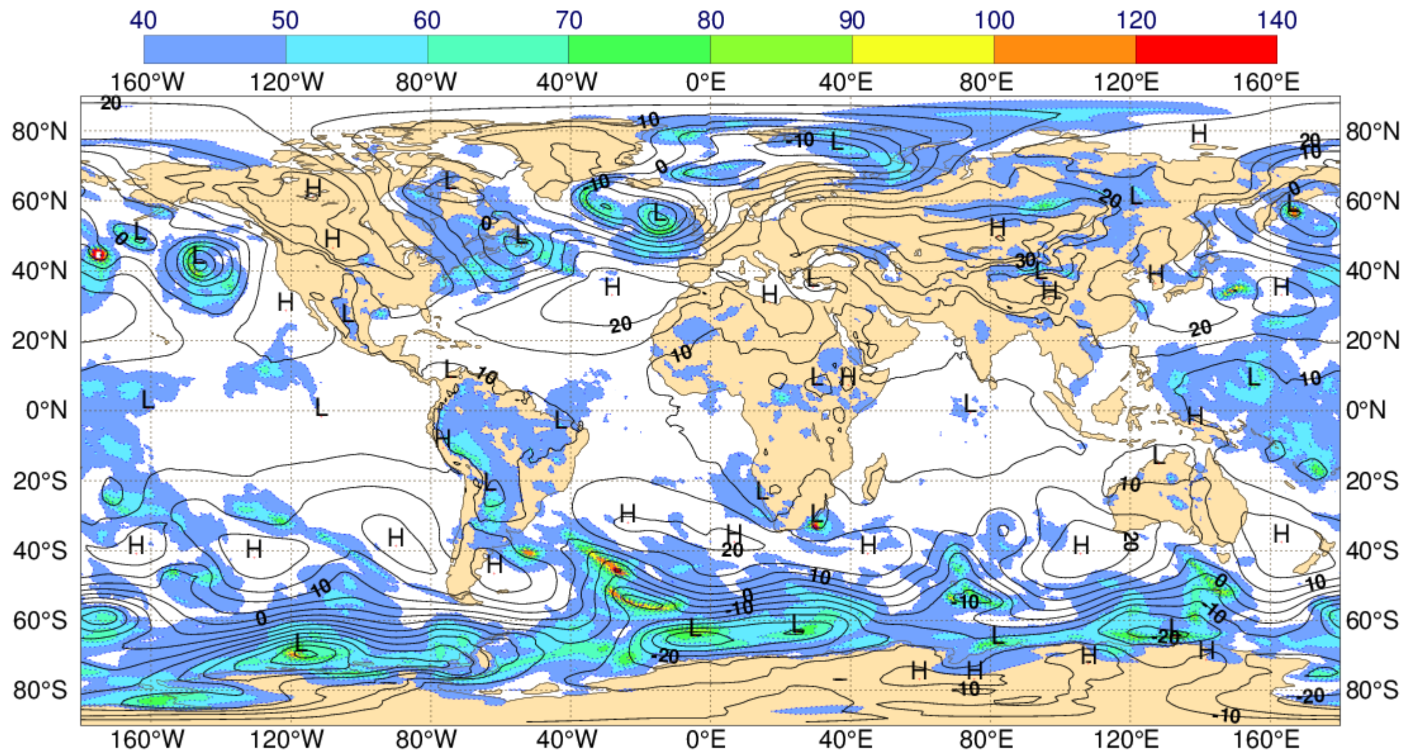
$$\mathbf{H}\mathbf{P}^b\mathbf{H}^T = \mathbf{H}\mathbf{X}^b(\mathbf{H}\mathbf{X}^b)^T =$$

$$\frac{1}{M-1}\Sigma_m(\mathbf{H}\mathbf{x}^b{}_m - <\mathbf{H}(\mathbf{x}^b{}_m)>)\ (\mathbf{H}\mathbf{x}^b{}_m - <\mathbf{H}(\mathbf{x}^b{}_m)>)^T$$

- Not having to code and maintain TL and ADJ operators is a distinct advantage!

Ensemble Kalman Filters

- In the EnKF the error covariances are sampled from the ensemble forecasts. They reflect the current state of the atmospheric flow



Standard deviation of surface pressure background t+6h fcst (shaded, Pa)
Z1000 background t+6h fcst (black isolines)

Ensemble Kalman Filters

- The Ensemble Kalman Filter is a Monte Carlo technique: it requires us to generate a sample $\{\mathbf{x}^b{}_m;\ m=1,..,M\}$ drawn from the pdf of background error: how to do this?

- We can generate this from a sample $\{\mathbf{x}^a{}_{t-1,m};\ m=1,..,M\}$ of the pdf of analysis error for the previous cycle:

$$\mathbf{x}^b{}_{t,m} = \mathcal{M}(\mathbf{x}^a{}_{t-1,m}) + \boldsymbol{\eta}_m$$

where $\boldsymbol{\eta}_m$ is a sample drawn from the pdf of model error.

- How do we generate a sample from the analysis pdf? Let us look at the analysis update again:

$$\mathbf{x}^a = \mathbf{x}^b + \mathbf{K}\,(\mathbf{y} - \mathbf{H}(\mathbf{x}^b)) = (\mathbf{I}\text{-}\mathbf{KH})\,\mathbf{x}^b + \mathbf{Ky}$$

- If we subtract the true state $\mathbf{x}^*$ from both sides (and assume $\mathbf{y}^*=\mathbf{Hx}^*$)

$$\mathbf{e}^a = (\mathbf{I}\text{-}\mathbf{KH})\,\mathbf{e}^b + \mathbf{Ke}^o$$

- i.e., the errors have the same update equation as the state

Ensemble Kalman Filters

- Consider now an ensemble of analysis where all the inputs to the analysis (i.e., the background forecast and the observations) have been perturbed according to their errors:

$$\mathbf{x}^{a\prime} = (\mathbf{I}\text{-}\mathbf{KH})\, \mathbf{x}^{b\prime} + \mathbf{K}\mathbf{y}'$$

- If we subtract the unperturbed analysis $\mathbf{x}^a = (\mathbf{I}\text{-}\mathbf{KH})\, \mathbf{x}^b + \mathbf{K}\mathbf{y}$

$$\boldsymbol{\varepsilon}^a = (\mathbf{I}\text{-}\mathbf{KH})\, \boldsymbol{\varepsilon}^b + \mathbf{K}\boldsymbol{\varepsilon}^o$$

- Note that the observations (during the update step) and the model (during the forecast step) are perturbed explicitly (i.e., we add random numbers with prescribed statistics).

- The background is implicitly perturbed , i.e.:

$$\mathbf{x}^b = \mathcal{M}(\mathbf{x}^a{}_{t-1,m}) + \boldsymbol{\eta}_m$$

- Hence, one way to generate a sample drawn from the pdf of analysis error is to perturb the observations and the model with perturbations drawn from their error covariances.

- The EnKF based on this idea is called Perturbed Observations (Stochastic) EnKF (Houtekamer and Mitchell, 1998). It is also the basis of ECMWF EDA (more on this later)

Ensemble Kalman Filters

- Another way to construct the analysis sample without perturbing the observations (but still perturb the model!) is to make a linear combination of the background sample:

$$\mathbf{X}^a = \mathbf{X}^b \mathbf{T}$$

where $\mathbf{T}$ is a $M \times M$ matrix chosen such that it produces the correct analysis covariance:

$$\mathbf{X}^a(\mathbf{X}^a)^T = (\mathbf{X}^b\mathbf{T})\,(\mathbf{X}^b\mathbf{T})^T = \mathbf{P}^a = (\mathbf{I}\text{-}\mathbf{KH})\mathbf{P}^b$$

- Note that the choice of $\mathbf{T}$ is not unique: Any orthonormal transformation $\mathbf{Q}$ ($\mathbf{Q}\mathbf{Q}^T = \mathbf{Q}^T\mathbf{Q} = \mathbf{I}$) can be applied to $\mathbf{T}$ and give another valid analysis sample
- Implementations also differ on the treatment of observations (i.e., local patches, one at a time)
- Consequently there are a number of different, functionally equivalent, implementations of the Deterministic EnKF (ETKF, Bishop *et al.*, 2001; LETKF, Ott *et al.*, 2004, Hunt *et al.*, 2007; EnSRF, Whitaker and Hamill, 2002; EnAF, Anderson, 2001;…)

Ensemble Kalman Filters

- We might want to ask the questions:
  1. How good is the EnKF for state estimation?
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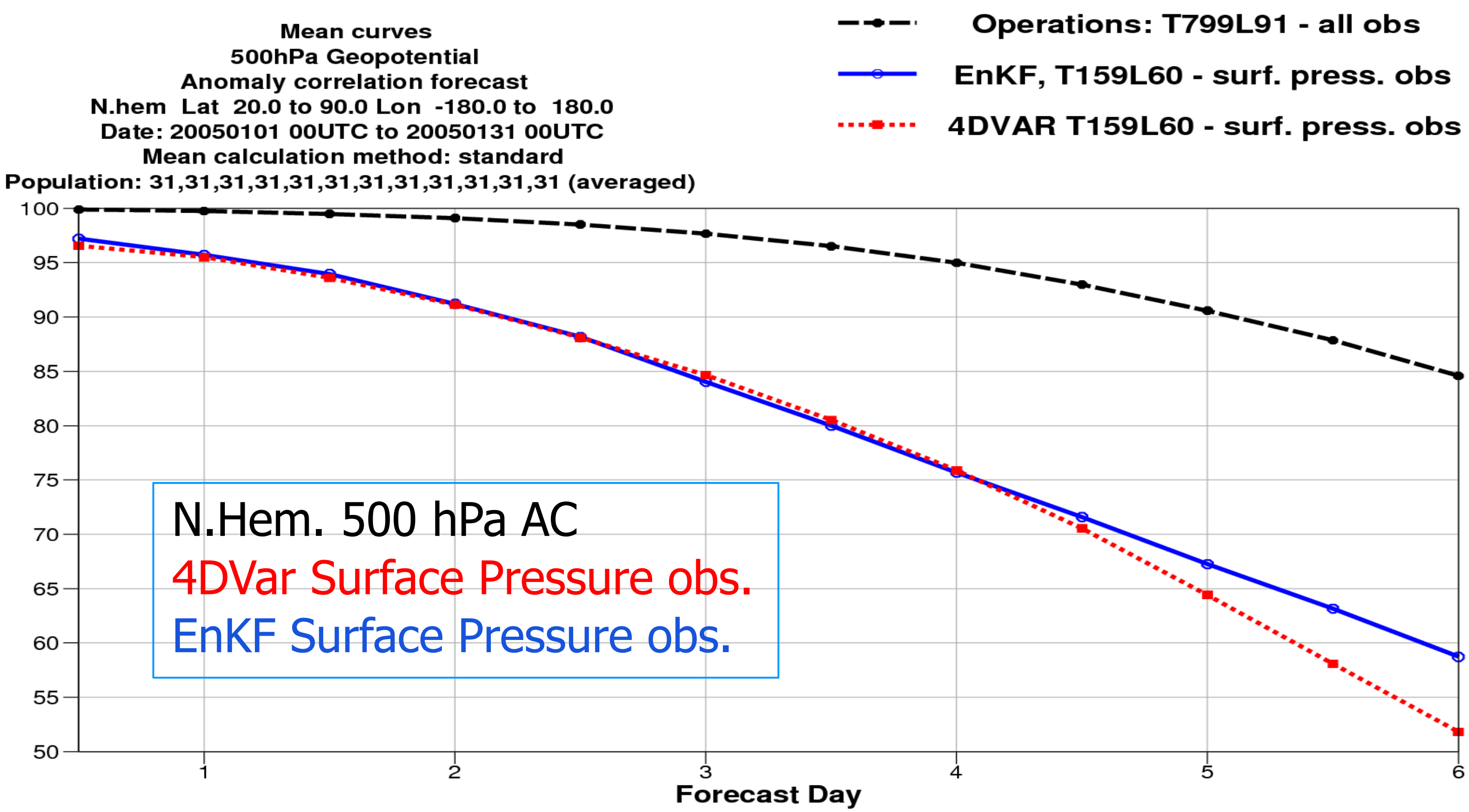
Ensemble Kalman Filters

- We might want to ask the questions:
  1. How good is the EnKF for state estimation?
  2. How does it compare with 4D-Var?

- The short answer: it depends…

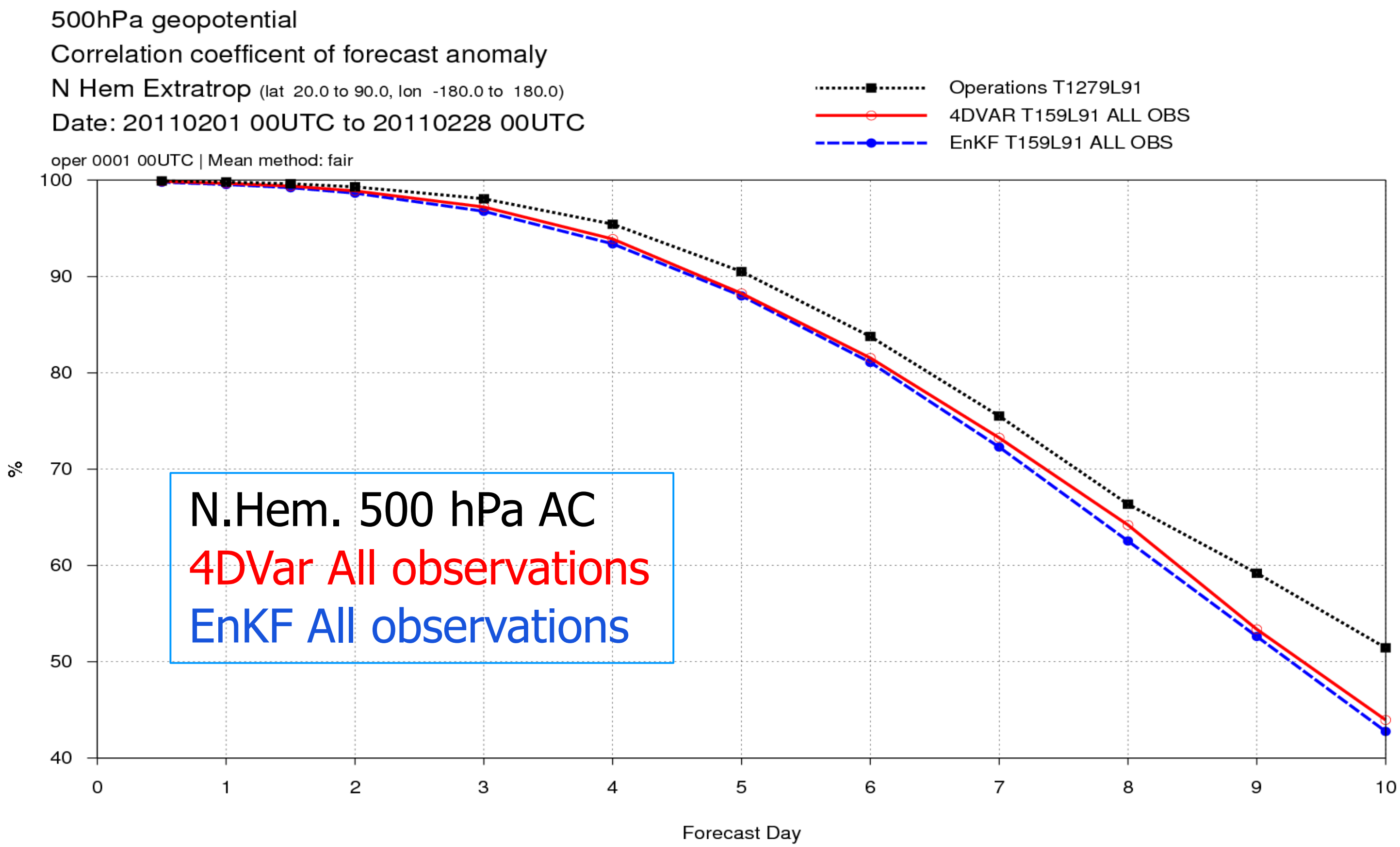
Ensemble Kalman Filters

- For sparsely observed systems the EnKF works quite well:



Ensemble Kalman Filters

- For fully observed systems the EnKF works not quite as well:



Ensemble Kalman Filters

- Pluses:
  1. Background error estimates reflect state of the flow
  2. Provides an ensemble of analyses: can use to initialise ensemble prediction
  3. Competitive with 4D-Var for sparsely observed systems
  4. Very good scalability properties
  5. Relative ease of coding and maintenance (No TL and ADJ models!)

Ensemble Kalman Filters

- Cons:

  1. The affordable ensembles are relatively small (O(100)), thus sampling noise and rank deficiency of the sampled error covariances become a performance limiting factor for the EnKF

  2. Careful localization of sampled covariances becomes necessary: This is an on-going research topic for both EnKF and Ensemble Variational hybrid systems

  3. Vertical covariance localization becomes conceptually and practically more difficult for observations (e.g., satellite radiances) which are non-local, i.e. they sample a layer of the atmosphere (Campbell *et al.*, 2010). There are ways around this (e.g., Lei et al, 2018 and references therein) but they are expensive, in that they require a much larger ensemble.

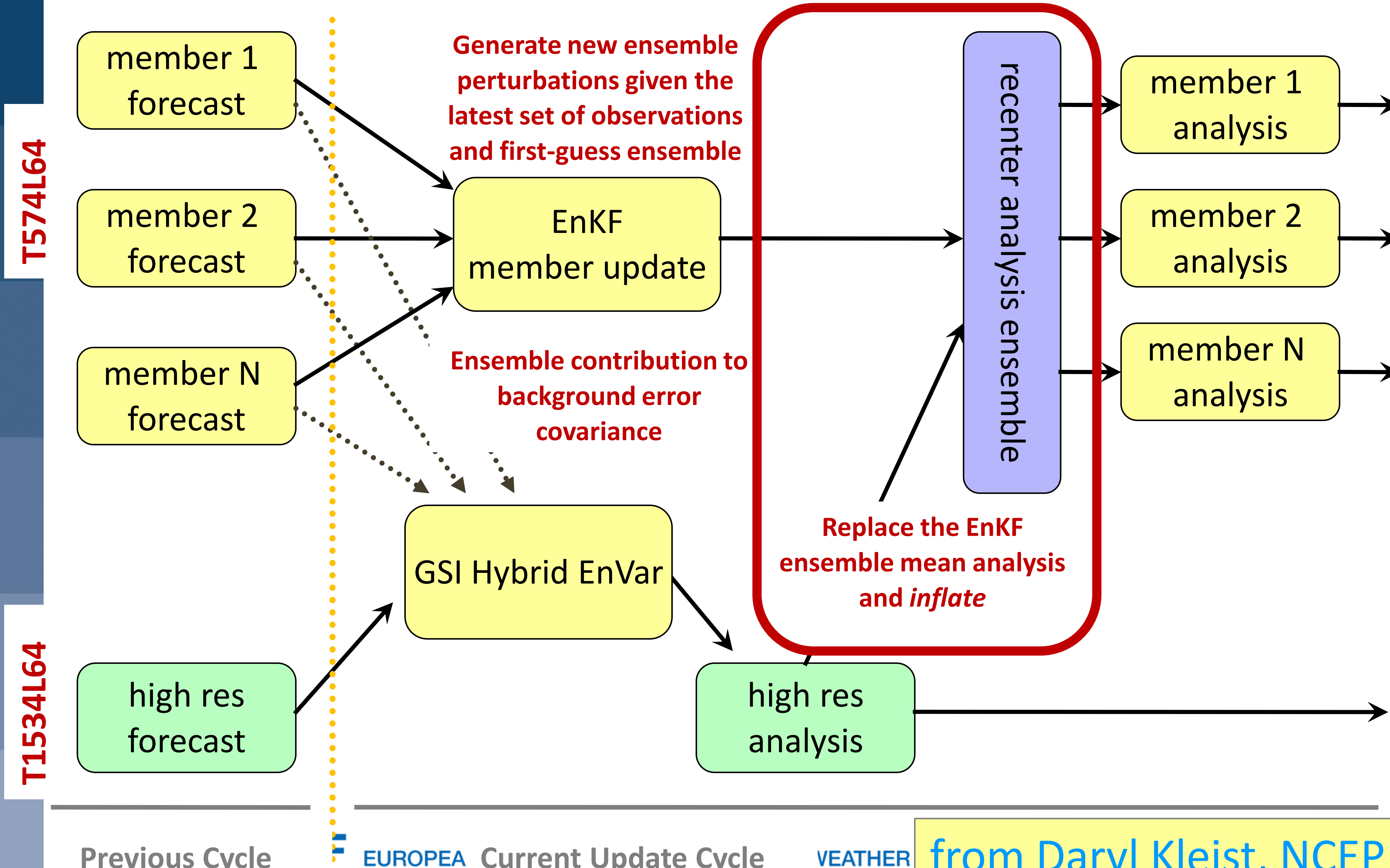
Ensemble Kalman Filters

- Cons:
  4. EnKF produces linear analysis updates. However high-resolution models and new cloud/precip. sensitive observations (e.g., rain radar, cloud/rain affected radiances, etc) are increasingly nonlinear. What to do?
     1. Iterate the analysis, similar to what incremental 4D-Var does (e.g., Iterative EnKF, Sakov, 2012; Sakov et al., 2018). Relatively easy but computationally expensive;
     2. Extend the Gaussian framework of the EnKF to classes of non-Gaussian pdfs. This can be done in a variety of ways, e.g. Gaussian Mixture models (Andersson and Andersson, 1999; Bengtsson et al. 2003; Hoteit et al. 2008, 2012; Stordal et al. 2011; Frei and Künsch, 2013), GIGGS Filter (Bishop, 2016);
     3. A combination of 1 and 2 (e.g., Posselt and Bishop, 2018);
     4. Rank Histogram Filters (Andersson, 2010; Metref et al, 2014)
     5. Employ some combination/hybrid of EnKF and Particle Filter (see for example: Van Leeuwen, Y. Cheng and S. Reich, 2015; Carrassi et al, 2017; Google this for even more recent results!)
     6. Let the EnKF give up gracefully in presence of increasing nonlinearity (Bonavita, Geer and Hamrud, 2019, in preparation)
     7. …

Ensemble Kalman Filters in hybrid DA

- While the pure EnKF is not currently competitive with variational methods for state estimation in global NWP, its good scalability properties and ease of maintenance make it a popular choice as a Monte Carlo system to estimate and cycle the error covariances ($P^{a/b}$) needed in a variational analysis system and to initialise an ensemble prediction system: hybrid Variational-EnKF analysis systems (NCEP, CMC, UKMO)

Dual-Res Coupled Hybrid Var/EnKF Cycling



from Daryl Kleist, NCEP

Summary

- For linear model $\mathbf{M}$ and the observation operators $\mathbf{H}$ the Kalman Filter produces an optimal (minimum error variance) sequence of analyses ($\mathbf{x}^a{}_1$, $\mathbf{x}^a{}_2$, …, $\mathbf{x}^a{}_{t-1}$, $\mathbf{x}^a{}_t$)
- We have shown that under the additional assumption of Gaussian errors the Kalman Filter provides the exact posterior probability estimate, $p(\mathbf{x}^a{}_t | \mathbf{x}^b{}_0; \mathbf{y}_0, \mathbf{y}_1, \ldots, \mathbf{y}_t)$.
- Kalman Filters are impractical for large-dimensional systems like in NWP, due to the impossibility of storing and evolving the state error covariance matrices ($\mathbf{P}^{a/b}$)
- We need to use reduced-rank representations of the state error covariance matrices: this can be done, but has other drawbacks (need for localisation, physical imbalances, etc.)
- The Ensemble Kalman Filter is a Monte Carlo implementation of the reduced-rank Kalman Filter. It works well for sparsely observed systems, but for well observed systems the severe rank reduction can be a difficult issue
- The EnKF (and its variants) are currently used in most global NWP Centres as the error cycling component of a hybrid Variational-EnKF system

Bibliography

1. Anderson, J. L., and S. L. Anderson, 1999: A Monte Carlo implementation of the nonlinear filtering problem to produce ensemble assimilations and forecasts. Mon. Wea. Rev., 127, 2741–2758, doi:https://doi.org/10.1175/1520-0493(1999)127<2741:AMCIOT>2.0.CO;2
2. Anderson, J. L., 2001. An ensemble adjustment Kalman filter for data assimilation. Mon. Wea. Rev. 129, 2884–2903.
3. Anderson, J. L.: A non-Gaussian ensemble filter update for dataassimilation, Mon. Weather Rev., 138, 4186–4198, 2010.
4. Bengtsson, T., C. Snyder, and D. Nychka, 2003: Toward a nonlinear ensemble filter for high-dimensional systems. J. Geophys. Res., 108, 8775, doi:https://doi.org/10.1029/2002JD002900.
5. Bishop, C. H., Etherton, B. J. and Majumdar, S. J., 2001. Adaptive sampling with ensemble transform Kalman filter. Part I: theoretical aspects. Mon. Wea. Rev. 129, 420–436.
6. Bishop, C.H., 2016: The GIGG-EnKF: ensemble Kalman filtering for highly skewed non-negative uncertainty distributions. Q. J. R. Meteorol. Soc., 142, 1395–1412.
7. Bishop, C. H., 2017: Gain Form of the Ensemble Transform Kalman Filter and Its Relevance to Satellite Data Assimilation with Model Space Ensemble Covariance Localization. Mon, Wea. Rev., https://doi.org/10.1175/MWR-D-17-0102.1
8. Bromiley, P.A., 2014: Products and Convolutions of Gaussian Probability Density Functions, TINA Memo N. 2003-003. Available at http://www.tina-vision.net/docs/memos/2003-003.pdf
9. Burgers, G., Van Leeuwen, P. J. and Evensen, G., 1998. On the analysis scheme in the ensemble Kalman filter. Mon. Wea. Rev. 126, 1719–1724.
10. Campbell, W. F., C. H. Bishop, and D. Hodyss, 2010: Vertical covariance localization for satellite radiances in ensemble Kalman Filters. Mon. Wea. Rev.,138,282–290.
11. Carrassi, A., Bocquet, M., Bertino, L., et al., 2017. Data assimilation in the geosciences—an overview on methods, issues and perspectives. arXiv:1709.02798.
12. Evensen, G., 1994. Sequential data assimilation with a nonlinear quasi-geostrophic model using Monte Carlo methods to forecast error statistics. J. Geophys. Res. 99(C5), 10 143–10 162.
13. Evensen, G . 2004 . Sampling strategies and square root analysis schemes for the EnKF . No.: 54 . p.: 539-560 . Ocean Dynamics

Bibliography

14. Fisher, M., Leutbecher, M. and Kelly, G. A. 2005. On the equivalence between Kalman smoothing and weak-constraint four-dimensional variational data assimilation. Q.J.R. Meteorol. Soc., 131: 3235–3246.
15. Flowerdew, J. 2015. Towards a theory of optimal localisation. Tellus A 2015, 67, 25257, http://dx.doi.org/10.3402/tellusa.v67.25257
16. Frei, M., and H. R. Künsch, 2013: Mixture ensemble Kalman filters. Comput. Stat. Data Anal., 58, 127–138, doi:https://doi.org/10.1016/j.csda.2011.04.013.
17. Hoteit, I., X. Luo, and D.-T. Pham, 2012: Particle Kalman filtering: A nonlinear Bayesian framework for ensemble Kalman filters. Mon. Wea. Rev., 140, 528–542, doi:https://doi.org/10.1175/2011MWR3640.1.
18. Houtekamer, P. L. and Mitchell, H. L., 1998. Data assimilation using an ensemble Kalman filter technique. Mon. Wea. Rev. 126, 796–811.
19. Houtekamer, P. L. and Mitchell, H. L., 2001. A sequential ensemble Kalman filter for atmospheric data assimilation. Mon. Wea. Rev. 129, 123–137.
20. Hunt, B. R., Kostelich, E. J. and Szunyogh, I., 2007. Efficient data assimilation for spatiotemporal chaos: a local ensemble transform Kalman filter. Physica D, 230, 112–126.
21. Lei, L., Whitaker, J. S., & Bishop, C. (2018). Improving assimilation of radiance observations by implementing model space localization in an ensemble Kalman filter. Journal of Advances in Modeling Earth Systems, 10, 3221–3232. https://doi.org/10.1029/2018MS001468
22. Liu C, Xiao Q, Wang B. 2008. An ensemble-based four-dimensional variational data assimilation scheme. part i: Technical formulation and preliminary test. Mon. Weather Rev. 136: 3363–3373.
23. Lorenc, A.C.,2003: The potential of the ensemble Kalman filter for NWP—A comparison with 4D-VAR. Q. J. R. Meteorol. Soc., 129: 3183–3203.
24. Ménétrier, B., T. Montmerle, Y. Michel and L. Berre, 2015: Linear Filtering of Sample Covariances for Ensemble-Based Data Assimilation. Part I: Optimality Criteria and Application to Variance Filtering and Covariance Localization. Mon. Wea. Rev, 143, 1622–1643

Bibliography

25. Metref, S., E. Cosme, C. Snyder, and P. Brasseur, 2014: A non-Gaussian analysis scheme using rank histograms for ensemble data assimilation. Nonlinear Processes Geophys., 21, 869–885, doi:https://doi.org/10.5194/npg-21-869-2014.
26. Miyoshi, T., K. Kondo, and T. Imamura, 2014: The 10,240-member ensemble Kalman filtering with an intermediate AGCM, Geophys. Res. Lett., 41, 5264–5271, doi:10.1002/2014GL060863.
27. Ott, E., Hunt, B. H., Szunyogh, I., Zimin, A. V., Kostelich, E. J. and co-authors. 2004. A local ensemble Kalman filter for atmospheric data assimilation. Tellus 56A, 415–428.
28. Periáñez, Á., Reich, H. and Potthast, R. 2014. Optimal localization for ensemble Kalman filter systems. J. Met. Soc. Japan. 62, 585 597.
29. Posselt DJ, Bishop CH. Nonlinear data assimilation for clouds and precipitation using a gamma inverse-gamma ensemble filter. Q J R Meteorol Soc. 2018;144:2331–2349. https://doi.org/10.1002/qj.3374
30. Sakov, P., D.S. Oliver, and L. Bertino, 2012: An Iterative EnKF for Strongly Nonlinear Systems. Mon. Wea. Rev., 140, 1988–2004, https://doi.org/10.1175/MWR-D-11-00176.1
31. Sakov, P., Haussaire, J.-M. and Bocquet, M., 2018: An iterative ensemble Kalman filter in the presence of additive model error. Q.J.R. Meteorol. Soc.. doi:10.1002/qj.3213
32. Stordal, A. S., H. A. Karlsen, G. Nævdal, H. J. Skaug, and B. Vallès, 2011: Bridging the ensemble Kalman filter and particle filters: The adaptive Gaussian mixture filter. Comput. Geosci., 15, 293–305, doi:https://doi.org/10.1007/s10596-010-9207-1.
33. Thépaut, J.-N., Courtier, P., Belaud, G. and Lemaître, G. 1996. Dynamical structure functions in a four-dimensional variational assimilation: A case-study. Q. J. R. Meteorol. Soc., 122, 535–561
34. van Leeuwen, P. J., Y. Cheng, and S. Reich,Nonlinear data assimilation, vol. 2, Springer, 2015
35. Whitaker, J. S. and Hamill, T. M., 2002. Ensemble data assimilation without perturbed observations. Mon. Wea. Rev. 130, 1913–1924.
36. Wikle, C.K. and Berliner, L.M., 2007: A Bayesian tutorial for data assimilation. Physica D: Nonlinear Phenomena, vol. 230, no. 1-2, pp. 1–16